Study of the $q^2$ dependence of $B \to \pi \ell \nu$ and of $B \to \rho(\omega) \ell \nu$ and extraction of $|V_{ub}|$

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We report on determinations of \(|V_{ub}|\) resulting from studies of the branching fraction and \(q^2\) distributions in exclusive semileptonic \(B\) decays that proceed via the \(b \to u\) transition. Our data set consists of the 9.7×10⁶ \(B\bar{B}\) meson pairs collected at the \(Y(4S)\) resonance with the CLEO II detector. We measure \(B(B^0 \to \pi^- \ell^+ \nu) = (1.33 \pm 0.18 \pm 0.11 \pm 0.01 \pm 0.07) \times 10^{-4}\) and \(B(B^0 \to \rho^- \ell^+ \nu) = (2.17 \pm 0.34 \pm 0.57 \pm 0.41 \pm 0.01) \times 10^{-4}\), where the errors are statistical, experimental systematic, systematic due to residual form-factor uncertainties in the signal, and systematic due to residual form-factor uncertainties in the cross-feed modes, respectively. We also find \(B(B^+ \to \pi^0 \ell^- \nu) = (0.84 \pm 0.31 \pm 0.16 \pm 0.09) \times 10^{-5}\), consistent with what is expected from the \(B \to \pi \ell \nu\) mode and quark model symmetries. We extract \(|V_{ub}|\) using light-cone sum rules for \(0 \leq q^2 \leq 16\) GeV² and lattice QCD for 16 GeV²\(\leq q^2 \leq q^2_{\text{max}}\). Combining both intervals yields \(|V_{ub}| = (3.24 \pm 0.22 \pm 0.13 \pm 0.09) \times 10^{-3}\) for \(\pi \ell \nu\), and \(|V_{ub}| = (3.00 \pm 0.21 \pm 0.35 \pm 0.28) \times 10^{-3}\) for \(\rho \ell \nu\), where the errors are statistical, experimental systematic, theoretical, and \(\rho \ell \nu\) form-factor shape, respectively. Our combined value from both decay modes is \(|V_{ub}| = (3.17 \pm 0.17 \pm 0.16 \pm 0.53 \pm 0.03) \times 10^{-3}\).

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I. INTRODUCTION

The element \(V_{ub}\) remains one of the most poorly constrained parameters of the Cabibbo-Kobayashi-Maskawa (CKM) matrix [1]. Its magnitude \(|V_{ub}|\) plays a central role in constraints based on the unitarity of the CKM matrix and inputs from both \(CP\)-conserving processes in the \(B\) meson decay and \(CP\)-violating processes in the neutral kaon and \(B\) systems. The value of \(|V_{ub}|\) and, in particular, the accuracy to which we have measured this important parameter, have been the subjects of considerable debate over the past decade [2]. An accurate determination of \(|V_{ub}|\) with well-understood uncertainties remains one of the fundamental priorities for heavy flavor physics.

A number of \(|V_{ub}|\) measurement approaches have been attempted, and are reviewed in Ref. [2]. Inclusive techniques are hampered by a mismatch in kinematic regions where the large experimental backgrounds from \(b \to c \ell \nu\) can be suppressed versus regions in which the theoretical uncertainties can be reliably determined. For exclusive reconstruction of particular final states, the primary challenge is the calculation of the form factors for those channels. The first measurements of exclusive charmless semileptonic branching fractions [3], including evaluation of \(|V_{ub}|\), were performed by the CLEO experiment at the Cornell Electron Storage Ring (CESR) using the modes \(B^0 \to \pi^- \ell^+ \nu\), \(B^+ \to \rho^0 \ell^+ \nu\), \(B^0 \to \rho^- \ell^+ \nu\), \(B^+ \to \omega \ell^+ \nu\), and charge-conjugate decays, where \(\ell = e\) or \(\mu\). A second measurement of the \(\rho \ell \nu\) modes by CLEO [4], using similar techniques but a much different signal to background optimization, provided consistent, essentially independent, results with a similar total uncertainty. The combined analyses yielded \(|V_{ub}| = (3.25 \pm 0.14 \pm 0.21 \pm 0.55) \times 10^{-3}\), where the errors are statistical, experimental systematic, and estimated theoretical uncertainties, respectively. The \(\pi\) and \(\rho\) modes contribute about equally to this result.

This paper presents an update of the original exclusive \(B \to X_u \ell \nu\) analysis [3], and is based on a total data sample of 9.7×10⁶ \(B\bar{B}\) pairs collected on the \(Y(4S)\) resonance. An additional data sample totalling 4.5 fb⁻¹ was collected off-resonance for the estimation of continuum backgrounds. The results presented here supersede those of Ref. [3]. In addition to using a larger data set, the analysis has been modified to minimize uncertainties arising from the momentum-transfer \((q^2)\) dependence of the form factors. Most notably, the lower bounds on the charged-lepton momentum for both the pseudoscalar and the vector modes have been lowered, and the branching fractions are determined independently in three \(q^2\) regions. For the \(\rho\) modes, the branching fractions as a function of \(q^2\) were first determined by the second CLEO \(\rho \ell \nu\) analysis [4]. The present analysis has a significantly broader accepted range for the charged lepton momentum, which allows for better discrimination among models. A detailed description of this analysis can be found in Ref. [5].

II. EXCLUSIVE CHARMLESS SEMILEPTONIC DECAYS

The semileptonic transition of a \(B\) meson (a pseudoscalar) to a final state with a single pseudoscalar meson \(P\) can, in the
limit of a massless charged lepton, be described by a single
form factor \( f_1(q^2) \):

\[
\frac{d\Gamma(B^0 \rightarrow P^- \ell^+ \nu)}{dyd\cos\theta_{\ell\nu}} = |V_{ub}|^2 \frac{G_F^2 k_P m_B^2}{32\pi^3} \sin^2\theta_{\ell\nu}|f_1(q^2)|^2,
\]

(1)

where \( y = q^2/M_B^2 \), \( M_B \) is the mass of the \( B \) meson, \( G_F \) is the Fermi constant, \( k_P \) is the meson momentum, and \( \theta_{\ell\nu} \) is the angle between the charged lepton direction in the virtual \( W \) \((\ell + \nu)\) rest frame and the direction of the virtual \( W \) in the \( B \) rest frame. For a transition to a final state with a single vector
meson \( V \), three form factors \( (A_1, A_2, \text{and } V) \) are necessary:

\[
\frac{d\Gamma(B^0 \rightarrow V^- \ell^+ \nu)}{dyd\cos\theta_{\ell\nu}} = |V_{ub}|^2 \frac{G_F^2 k_V M_B^2}{128\pi^3} \left[ (1 - \cos\theta_{\ell\nu})^2 |H_+|^2 / 2 
\right.

\[
+ (1 + \cos\theta_{\ell\nu})^2 |H_-|^2 / 2 
\]

\[
+ \sin^2\theta_{\ell\nu}|H_0|^2 \right],
\]

(2)

where \( k_V \) is the meson momentum and the three helicity
amplitudes are given by

\[
H_\pm = \frac{1}{M_B + m_V} [ A_1(q^2) \mp 2M_B k_V V(q^2) ],
\]

(3)

and

\[
H_0 = \frac{1}{\sqrt{2}} \frac{M_B}{2m_V(M_B + m_V)} \times \left[ \left( 1 - \frac{m_V^2}{M_B^2} - y \right) A_1(q^2) - 4k_V^2 A_2(q^2) \right].
\]

(4)

The structure of these differential decay rates immediately
allows us to draw some general conclusions regarding the
properties of the semileptonic decays that we reconstruct in
this analysis. For the \( p(\omega)\ell\nu \) transitions, the left-handed,
\( V^-A \), nature of the charged current at the quark level mani-

fests itself at the hadronic level as \( |H_-| > |H_+| \). The \( H_- \) con-
tribution is also expected to dominate the \( H_0 \) con-
tribution, leading to a forward-peaked distribution for \( \cos\theta_{\ell\nu} \).

The pseudoscalar modes also contain an extra factor of the meson
momentum squared, which suppresses the rate near \( q^2_{\text{max}} \) \((k_P = 0)\). Taken together, these two
effects give the pseudoscalar modes a softer charged lepton
momentum spectrum than the vector modes.

Calculation of the form factors has become a considerable
theoretical industry, with a variety of techniques now being
employed. Form factors based on lattice QCD (LQCD) cal-
culations \([6-18] \) and on light-cone sum rules (LCSR) \([19-27] \) currently have uncertainties in the 15% to 20% range. A
variety of quark-model calculations exist \([28-42] \). Finally, a
number of other approaches \([43-48] \), such as dispersive

bounds and experimentally constrained models based on
heavy quark symmetry, all seek to improve the range of \( q^2 \)
over which the form factors can be estimated without intro-
duction of significant model dependence. Figure 1 illustrates
the broad variation in shape that arises in the literature. Un-
fortunately, all the form-factor calculations currently have
contributions to the uncertainty that are uncontrolled. The
light-cone sum rules calculations assume quark-hadron dual-
ity, offering a “canonical” contribution to the uncertainty of
10%, but with no known means of rigorously estimating that
uncertainty. The LQCD calculations to date remain in the
“quenched” approximation (no light quark loops in the
propagators), which limits the ultimate precision to the 15% to
20% range. With the quark-model calculations it is diffi-
cult to quantify the uncertainty of a particular calculation by
their very nature. These uncertainties in the form factors
translate directly into the same fractional uncertainty on
\( |V_{ub}| \).

In the \( \pi\ell\nu \) modes, with only a single form factor in the
massless lepton approximation, we expect that the rates ex-
tracted in the \( q^2 \) intervals that we have chosen will be largely
independent of the form-factor shapes. In the vector modes,
however, the three form factors interfere and differences in
this interference among models, particularly at lower \( q^2 \)
values, can lead to a residual model dependence. To investigate
this effect, we will analyze the vector modes with three separate
charged lepton momentum requirements.

**III. EVENT RECONSTRUCTION AND SELECTION**

The CLEO detector \([49,50] \) contains three concentric
tracking devices within a 1.5 T superconducting solenoid that
detect charged particles over 95% (93%) of the solid
angle for the first third (last two-thirds) of the data. For the last
two-thirds of the data, a silicon vertex detector replaced a
straw-tube wire chamber. The momentum resolution at
2 GeV/c is 0.6%. A CsI(Tl) electromagnetic calorimeter,
also inside the solenoid, covers 98% of 4 \( \pi \). A typical \( \pi^0 \)
mass resolution is 6 MeV. Charged tracks are assigned the
most probable mass based on specific ionization, time of
flight, and the relative rates as a function of momentum for
proton, \( K^+ \), and \( \pi^+ \) production in \( B \) decay.

The undetected neutrino complicates analysis of semilep-
tonic decays. Because of the good Hermiticity of the CLEO detector, we can reconstruct the neutrino via the missing energy (\(E_{\text{miss}} = 2E_{\text{beam}} - \Sigma E_i\)) and missing momentum (\(\vec{P}_{\text{miss}} = -\Sigma \vec{p}_i\)) in each event. In the process \(e^+e^- \rightarrow Y(4S) \rightarrow B\bar{B}\), the total energy of the beams is imparted to the \(B\bar{B}\) system; at CESR, that system is at, or nearly at, rest. (A small crossing angle has been in use at CESR for most of the running.) The missing mass, \(M^2_{\text{miss}} = E_{\text{miss}}^2 - |\vec{P}_{\text{miss}}|^2\), must be consistent, within resolution, with a massless neutrino. Specifically, we require \(-0.5 < M^2_{\text{miss}}/2E_{\text{miss}} < 0.3\) GeV for events with a total charge \(\Delta Q = 0\), and \(|M^2_{\text{miss}}|/2E_{\text{miss}} < 0.3\) GeV for events with \(|\Delta Q | = 1\).

Signal Monte Carlo (MC) events show a \(\vec{P}_{\text{miss}}\) resolution of 85 MeV/c. The resolution on \(E_{\text{miss}}\) is about three times larger than the momentum resolution [51]. Significant effort has been devoted to minimizing multiple counting of charged particles in the track reconstruction (e.g., particles that curl multiple times within the tracking volume), and to suppressing clusters in the calorimeter from charged hadrons that have interacted.

With an estimate of the neutrino four-momentum in hand, we can employ full reconstruction of our signal modes. Because the resolution on \(E_{\text{miss}}\) is so much larger than that for \(|\vec{P}_{\text{miss}}|\), we use \((E_{\ell v}, \vec{P}_{\ell v}) = (|\vec{P}_{\text{miss}}|, \vec{P}_{\text{miss}})\) for full reconstruction. The neutrino combined with the signal charged lepton (\(\ell\)) and meson (\(m\)) should satisfy, within resolution, the constraints on energy, \(\Delta E = (E_{\ell +} + E_{\ell -} + E_m) - E_{\text{beam}} = 0\), and on momentum, \(M^2_{\ell m v} = [E^2_{\text{beam}} - (\vec{p}_{\ell +} + \vec{p}_{\ell -} + \vec{p}_m)^2]^{1/2} = M_B\). The neutrino momentum resolution dominates the \(\Delta E\) resolution, so the momentum scaling corrects for the mismeasurement of the magnitude of the neutrino momentum in the \(M_{\ell m v}\) calculation. Uncertainty in the neutrino direction then remains as the dominant source of smearing in this mass calculation.

We reconstruct \(q^2 = M^2_{\ell m v} = (p_{\ell +} + p_{\ell -})^2\) for each decay from the reconstructed charged lepton four-momentum and the missing momentum. In addition to using the scaled reconstructed momentum \(\alpha \vec{p}_m\) described above, the direction of the missing momentum is changed through the smallest angle consistent with forcing \(M^2_{\ell m v} = M_B^2\). This procedure results in a \(q^2\) resolution of 0.3 GeV\(^2\), independent of \(q^2\). The \(\pi\ell\nu\) and the \(\rho\ell\nu\) modes are analyzed separately in the intervals \(q^2 < 8\) GeV\(^2\), \(8 < q^2 < 16\) GeV\(^2\), and \(q^2 > 16\) GeV\(^2\). For the \(a\ell\nu\) and \(g\ell\nu\) modes, for which we have low statistics, we sum over all \(q^2\).

Information from specific ionization is combined with calorimetric and tracking measurements to identify electrons with \(p_\ell > 600\) MeV/c over 90% of the solid angle. Particles registering hits in counters deeper than 5 interaction lengths over the polar angle range \(|\cos \theta| < 0.85\) are considered muons. Those with hits beyond 3 interaction lengths over \(|\cos \theta| < 0.71\) are used in a multiple-lepton veto, described below. Candidate leptons must have \(p_\ell > 1.0\) GeV/c for the \(\pi\) and \(\eta\) (pseudoscalar) modes, and \(p_\ell > 1.5\) GeV/c for the \(\rho\) and \(\omega\) (vector) modes, which can couple to the \(W\) helicities \(\pm 1\) and hence have a harder spectrum. This momentum requirement for the vector modes defines the nominal analysis. We also analyze the vector modes with the lepton momentum requirements \(p_\ell > 1.75\) GeV/c and \(p_\ell > 2.0\) GeV/c. The identification efficiency above 1.5 GeV/c averages over 90%; the probability that a hadron is misidentified as an electron (muon), a fake lepton, is about 0.1% (1%).

The 5-interaction-length requirement for muons causes the muon acceptance to fall rapidly below 1.4 GeV/c. As a result, only electrons contribute at the low end of the momentum range we accept for \(\pi\ell\nu\), and electrons dominate the measurement in the lowest \(q^2\) interval.

A \(\pi^0\) candidate must have a \(\gamma\gamma\) mass within 2 standard deviations of the \(\pi^0\) mass. We reconstruct the \(\omega\) via its \(\pi^+\pi^-\pi^0\) decay, reducing combinatoric background by rejecting combinations away from the center of the \(\omega\) Dalitz plot. In particular, we require \(|a|^2/a_{\text{max}}^2 > 0.4\), where \(a\) is the decay amplitude for the reconstructed point in the Dalitz plot, and \(a_{\text{max}}\) is the maximum amplitude at the center of the Dalitz plot. We reconstruct \(\eta\) in both the \(\gamma\gamma\) and the \(\pi^+\pi^-\pi^0\) decay modes. For the \(\gamma\gamma\), we require the reconstructed mass to be within 2 standard deviations of the \(\eta\) mass (within about 26 MeV). For the \(\pi^+\pi^-\pi^0\), we require \(|m_{\pi^+\pi^-\pi^0} - m_\eta| < 10\) MeV (about 1.7 times the resolution). We impose a kinematic mass constraint on the momentum of all \(\pi^0\) or \(\eta\) candidates in the \(\gamma\gamma\) final state.

Backgrounds arise from the \(e^+e^- \rightarrow q\bar{q}\) and \(e^+e^- \rightarrow \pi^+\pi^-\) continuum, fake leptons, \(b\rightarrow c\ell\nu\), and \(B\rightarrow X_s\ell\nu\) modes other than the signal modes. Backgrounds from continuum processes are suppressed by use of two event-shape variables. The selection criteria were optimized using background and signal Monte Carlo samples, rather than data, to avoid potential bias. The first variable is the angle \((\cos \theta_{\text{thruss}})\) between the thrust axis evaluated for the candidate signal-mode particles (not including the neutrino) and that for the rest of the event. (The thrust axes are signed by picking the hemisphere containing the most energy.) For \(B\bar{B}\) events at CESR, the distribution in this variable is flat because the \(B\)'s are nearly at rest and thus their decay orientations are independent. For continuum events the distribution is strongly forward and backward peaked. The ratio \(R_2\) of the second to the zeroth Fox-Wolfram moment [52], which distinguishes spherical from jetty topologies, is also utilized. The continuum background tends to have a small reconstructed \(q^2\). We therefore tune the continuum cut employed in the \(R_2\) plane separately in each \(q^2\) interval, and separately for the \(\pi\) and \(\rho\) modes. Signal events with low \(q^2\) appear rather jetty, so a cut using \(R_2\), when data are binned over a broad \(q^2\) range, would introduce an efficiency bias. So for the \(\omega\) and \(\eta\) modes, for which all \(q^2\) regions are combined, only a \(\cos \theta_{\text{thruss}}\) cut is applied, reducing uncertainties from the \(q^2\) dependence of the form factors. Our criteria suppress the continuum background by over a factor of 10 and are about 80% efficient.

The \(|p_\ell|\) cuts greatly reduce background from \(b\rightarrow c\rightarrow s\ell\nu\) and bias mildly against \(b\rightarrow c\ell\nu\). For the vector modes, we further require \(\cos \theta_{\text{thruss}} > 0\), since the signal rate is largely suppressed by \(V\rightarrow A\) outside this region, while the

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background is roughly flat in the region excluded, and falls off in the region accepted.

Backgrounds, particularly $b \to c \ell \nu$, can smear into the signal region in $\Delta E$ and $M_{m\ell\nu}$ when $p_{\text{miss}}$ misrepresents $p_\nu$. Such backgrounds are highly suppressed by rejecting events with multiple charged leptons or a total event charge $|\Delta Q| > 1$, both of which indicate missing particles. Requiring $M_{m\ell\nu}^2$ to be consistent with zero also provides powerful background suppression. Still, Monte Carlo studies show that the dominant remaining $b \to c \ell \nu$ events contain either a $K_L$ meson or a second neutrino (from $c \to s \ell \nu$, with the lepton not identified) that is roughly collinear with the primary neutrino.

Our selection criteria studies, based on statistical considerations, indicated that keeping the $|\Delta Q| = 1$ sample as well as the $\Delta Q = 0$ was favorable in spite of the poorer signal-to-background ratio. Further systematic considerations indicated that the use of the $|\Delta Q| = 1$ sample remained advantageous for the pseudoscalar modes. For the vector, in particular the $\rho$ modes, however, the overall poorer signal-to-background ratio made the $|\Delta Q| = 1$ sample overly sensitive to systematic effects in both the modelling of the $B \to X_u \ell \nu$ backgrounds and the simulation of the detector. Therefore for the vector modes we require $\Delta Q = 0$.

IV. EXTRACTION OF BRANCHING FRACTIONS

A. Method and binning

To extract the branching fraction information, we performed a binned maximum likelihood fit that was extended to include the finite statistics of the Monte Carlo, off-resonance, and fake-lepton samples following the method of Barlow and Beeston [59]. The data in each mode were coarsely binned over the two-dimensional region $5.175 \leq M_{m\ell\nu} \leq 5.2875$ GeV, $|\Delta E| < 0.75$ GeV. Further we binned both the data in the closest $2\pi$ and $3\pi$ masses in the $\rho$ and $\omega$ modes. The $|\Delta Q| = 1$ samples were binned separately from $\Delta Q = 0$ samples. Separation of the net charge samples allowed us to take advantage of the better signal-to-noise ratio of the $|\Delta Q| = 0$ sample while reducing our dependence on our knowledge of the absolute tracking efficiency. Finally, we binned the data in $q^2$ for the two $\pi\ell\nu$ and the two $\rho\ell\nu$ modes. For the $\omega \ell \nu$ and the $\eta \ell \nu$ modes, we combined all $q^2$ information into a single bin.

Our fitting strategy was designed to minimize dependence of the results on the details of the simulation—both from detector and physics standpoints. The choice of binning balanced separation of signal and background against reliance on detailed MC shape predictions. To help minimize the model dependence of the branching fraction determinations, we did not use information from the lepton momentum spectrum or from $\cos \theta_{\ell\ell}$ within the fit. Extraction of rates in the separate $q^2$ intervals further reduces reliance on the form factors.

The $\Delta E$ bin intervals used in the nominal fit were

$$-0.75 \leq \Delta E < -0.45 \text{ GeV}, \quad -0.45 \leq \Delta E < -0.15 \text{ GeV}, \quad \text{and} \quad -0.15 \leq \Delta E < 0.25 \text{ GeV}$$

(3) The $\Delta E$ signal band. The $M_{m\ell\nu}$ bin intervals were $5.175 \leq M_{m\ell\nu} < 5.2425$ GeV and 5.2425 $\leq M_{m\ell\nu} < 5.2875$ GeV. In the $\Delta E$ signal band, this second mass interval is divided into two equal bands. Hence we used a total of seven bins in these two variables. In the $\rho \ell \nu$ ($\omega \ell \nu$) modes, we used three equal bins over the $2\pi$ ($3\pi$) mass range within $\pm 285$ MeV ($\pm 30$ MeV) of the nominal $\rho$ ($\omega$) mass. The three $q^2$ intervals in the $\pi\ell\nu$ and the $\rho\ell\nu$ modes were $q^2 < 8$ GeV$^2$, $8 < q^2 < 16$ GeV$^2$, and $q^2 \geq 16$ GeV$^2$. The number of bins for each mode in the nominal fit is summarized in Table I. The nominal fit had a total of 259 bins. For studies in which the $|\Delta Q| = 1$ sample is included in the $\rho$ and $\omega$ modes, the fit had an additional 147 bins for a total of 406 bins.

To examine yields, efficiency, and kinematics in this paper, we use the most sensitive bin (the “signal bin”) $5.265 \leq M_{m\ell\nu} < 5.2875$ GeV and $-0.15 \leq \Delta E < 0.25$ GeV, though neighboring bins also contribute information to the fit. For comparison, the $M_{m\ell\nu}$ and $\Delta E$ resolutions are about 7 MeV and 100 MeV, respectively, dominated by the resolution on $|p_\nu|$. The $2\pi$ ($3\pi$) mass intervals $\pm 95$ MeV and $\pm 10$ MeV, centered on the nominal masses, are used for figures for $\rho$ and $\omega$ candidates, respectively.

To simplify the statistical interpretation of the results, we limited the number of multiple entries per event. For each individual mode, the candidate with the smallest $|\Delta E|$ among those satisfying $M_{m\ell\nu} > 5.175$ GeV was chosen, independent of $q^2$. A given event could contribute to multiple modes, although contribution near the signal region in more than one mode was rare. In the $\rho$ and $\omega$ modes, each of the mass bins described above was considered a separate mode.

B. Fit components and parameters

MC simulation provided the distributions in each mode for signal, the $b \to c$ background, the cross-feed among the modes, and the feed down from higher mass $B \to X_s \ell\nu$ decays. It included a full description of the $b \to c$ and charm decay modes and a GEANT-based [53] detector model. The $X_s \ell\nu$ feed down was evaluated with a simulation of the $B \to X_s \ell\nu$ process based on an inclusive operator product expansion (OPE) calculation [54] of $d\Gamma(B \to X_s \ell\nu)/dM_{X_s}$ using parameters determined from the CLEO analysis of the $B \to X_s \gamma$ photon spectrum [55,56] (also used in the recent CLEO lepton-momentum end-point analysis [57]). The nominal analysis combined this inclusive spectrum with the ISGW II model [31] for all mesons through the $\rho(1450)$. For
each exclusive mode, we “subtracted rate” from the inclusive calculation with a weight of the form $\exp[-\alpha(M_{X_{\ell\nu}} - M_R)^2/\Lambda_{QCD}^2]$, where $M_R$ is the central mass of the resonance $R$. At any given $M_{X_{\ell\nu}}$, the rate remaining after this subtraction of the exclusive modes is hadronized nonresonantly. Variations of the inclusive parameters based on the uncertainties in the $B\rightarrow X_{\gamma}$ analysis and variations of the hadronization model (e.g., fully nonresonant but with $\pi\pi\ell\nu$ removed from the $\rho$ mass region) are included in the systematic uncertainties. The signal modes are excluded from these $B\rightarrow X_{\gamma}\ell\nu$ samples.

The contributions from events in which hadrons have faked the signal leptons and from continuum are evaluated using data. The electron and muon identification fake rates from pions, kaons, and protons are measured in data using a variety of tagged samples. The analysis is performed on a sample of hadronic events with no identified leptons, treating each track in turn as a signal electron and then a signal muon. The contribution in each mode is weighted according to the fake rate.

We determined the residual continuum background using data collected 60 MeV below the $Y(4S)$ energy. The center-of-mass energy and cross-section differences were taken into account as necessary. For each combination of mode, reconstructed $q^2$ bin, and for each $\Delta Q$ value, we determined the rate over the full $\Delta E\cdot M_{m\ell\nu}$ plane by applying all cuts, including continuum-suppression cuts, and then scaling according to the relative on-resonance and off-resonance luminosities. To smooth the statistical fluctuations within each combination, we determined the shape over the $\Delta E\cdot M_{m\ell\nu}$ plane by applying the procedure. First, we dropped the continuum-suppression cuts and obtained the shape over the $\Delta E\cdot M_{m\ell\nu}$ plane for each combination from data. Then, from continuum $q\bar{q}$ MC calculations, $\tau^+\tau^-$ MC, and our faked lepton samples, we determined the change in shape over the $\Delta E\cdot M_{m\ell\nu}$ plane caused by application of the continuum-suppression cuts, i.e., we obtained the ratio of yields, with to without cuts, for each $\Delta E\cdot M_{m\ell\nu}$ bin, for each combination. Within the MC statistics (about 4.5 times the on-resonance continuum contribution in data), the predicted ratios were consistent with flat (no change in shape). Applying the ratios so obtained to the off-resonance data without continuum-suppression cuts, we obtained the shape of the background over the $\Delta E\cdot M_{m\ell\nu}$ plane, for each combination. The overall normalization for each combination was, again, determined from the observed rate in the off-resonance data for that combination.

For each signal mode, we generated a sample of signal Monte Carlo that is flat in phase space and processed these samples with our GEANT-based detector simulation. As we analyze each reconstructed event, we reweight the event to correspond to a particular calculation for the form factors involved in the decay. This procedure allowed us to sample a variety of form factor calculations. For each mode, we determine the efficiency matrix for reconstructed versus true $q^2$. Given our resolution and binning, the matrix is essentially diagonal, as Table II shows for the $\pi\ell\nu$ form-factor calculation of Ball and Zwicky (Ball’01) [27].

<table>
<thead>
<tr>
<th>True $q^2$ (GeV$^2$)</th>
<th>Reconstructed $q^2$</th>
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<tr>
<td>0–8</td>
<td>0.07, 0.001</td>
</tr>
<tr>
<td>8–16</td>
<td>4.4, 0.06</td>
</tr>
<tr>
<td>$\geq$16</td>
<td>4.4, 0.00</td>
</tr>
</tbody>
</table>

For these results, we have examined the following form factors for the signal modes and cross-feed rates. For $\pi\ell\nu$: Ball and Zwicky (light-cone sum rules) [27], ISGW II (a nonrelativistic quark model) [31], and the skewed parton distributions (SPD) of Feldmann and Kroll [39]. Other LQCD and LCSR calculations are also considered in extracting $|V_{ud}|$. For $\rho\ell\nu$: Ball and Braun (light-cone sum rules—Ball’98) [20], ISGW II, Melikhov and Stech (a relativistic quark model—Melikhov’00) [38], and UKQCD (a LQCD calculation—UKQCD’98) [8]. For $\eta\ell\nu$, we have only considered the ISGW II form factor. The above choices for $\pi\ell\nu$ and $\rho\ell\nu$ bracket the extremes in the variation of the shape of $d\Gamma/dq^2$ and hence provide a conservative estimate of the theoretical uncertainty on the branching fractions. In general, the theory references provide minimal guidance on the theoretical uncertainty in the form-factor shapes, and the variation among the chosen calculations appears larger than the variation expected within a given calculation. For nominal yields and figures, we use Ball’01 for the $\pi$ modes and Ball’98 for the vector modes.

We fit all the signal modes simultaneously. The parameters for the three $\pi^-\ell^+\nu$ $q^2$ intervals, the three $\rho^-\ell^+\nu$ $q^2$ intervals, and the total $\eta\ell\nu$ branching fraction floated as free parameters in the fit, for a total of 7 signal parameters. The isospin and quark symmetry relations $\Gamma(B^0\rightarrow\pi^0\ell\nu) = 2\Gamma(B^+\rightarrow\pi^+\ell\nu)$ and $\Gamma(B^0\rightarrow\rho^0\ell\nu) = 2\Gamma(B^+\rightarrow\rho^+\ell\nu)$, and $\Gamma(B^+\rightarrow\omega\ell\nu) = 2\Gamma(B^+\rightarrow\pi^0\ell\nu)$ constrain the rates for $B^+$ relative to $B^0$, and are assumed to hold for each $q^2$ region. We combined the three $\omega\ell\nu$ rate predictions that result from the quark symmetry assumption and the three $\rho\ell\nu$ rates to obtain the fit prediction for the total observed reconstructed $\omega\ell\nu$ yield. As mentioned above, only this integrated yield for $\omega\ell\nu$ contributes to the likelihood. The two $\eta$ submodes are tied to the total $\eta\ell\nu$ branching fraction by the measured $\eta$ branching fractions and the submode reconstruction efficiencies. To implement the isospin constraints, we assume equal charged and neutral $B$ production, $f_{+-} = f_{00}$, and input a lifetime ratio of $1.083\pm0.017$ [58]. For self-consistency, the cross-feed rates are constrained to the observed yields.

The $b\rightarrow c$ normalization in the fit varies independently for each mode, and within each mode for $\Delta Q = 0$ and $|\Delta Q| = 1$. The normalizations obtained are generally within 10% of those derived from luminosity and cross sections. The nominal fit therefore has an additional 11 free parameters for these normalizations.
We float the overall normalization of the generic \( B \to X_u \ell \nu \) feed-down sample, determining it from the fit. To help in determining that normalization, we take advantage of CLEO’s recent measurement [57] of the branching fraction for \( b \to u \ell \nu \) decays with leptons in the 2.2–2.6 GeV/c momentum range: \( \mathcal{B}(B \to X_u \ell \nu, 2.2 \leq p_\ell \leq 2.6 \text{ GeV/c}) = (2.30 \pm 0.38) \times 10^{-4} \) (the “end-point branching fraction’’). We constrained the \( B \to X_u \ell \nu \) feed-down normalization by adding a \( \chi^2 \) term to the log likelihood of the fit:

\[
-2 \ln L \to -2 \ln L + \frac{(B_{em} - B_{cp})^2}{\sigma_{em}^2},
\]

where \( B_{em} \) is the measured end-point branching fraction, \( \sigma_{em} \) is the total experimental uncertainty on that measurement, and \( B_{cp} \) is the branching fraction implied by the fit parameters. The fit prediction in each iteration is given by

\[
B_{cp} = B_{u \ell \nu} f_{u \ell \nu} + \sum_i \sum_m B_{m,i} f_{m,i}, \quad (m) \in \{ \pi^+, \pi^0, \rho^+, \rho^0, \omega, \eta \},
\]

where \( m \in \{ \pi^+, \pi^0, \rho^+, \rho^0, \omega, \eta \} \), \( B_{m,i} \) is the branching fraction for the decay mode \( m \) and the \( q^2 \) interval \( i \) in that iteration, \( f_{m,i} \) is the fraction of charged leptons, for that mode and \( q^2 \) interval, that is predicted by the form-factor calculation to lie in the end-point region, \( B_{u \ell \nu} \) is the branching fraction for the \( B \to X_u \ell \nu \) feed-down background in that iteration, and \( f_{u \ell \nu} \) is the fraction of charged leptons in the end-point momentum range obtained from our model.

The systematic error evaluation for the \( B \to X_u \ell \nu \) feed down, and checks using alternative procedures, are described below. The normalization is floated independently for each systematic variation of the various Monte Carlo, continuum, or fake samples described below so that the effect on the background normalization of mismodeling within the simulation is properly assessed.

In summary, we have 19 free parameters in the fit: the seven signal rates, the 11 generic \( b \to c \) background normalizations, and the one generic \( B \to X_u \ell \nu \) feed-down background normalization. The continuum background and fake-lepton background samples are absolutely normalized and their rates do not float in the fit. In fits discussed below for which we include the \( |\Delta Q| = 1 \) information in the vector-meson modes, there are an additional 3 \( b \to c \) background normalization parameters, for a total of 22 free parameters.

C. Checks and results

We have examined the reliability of our fitting procedure via a bootstrap technique. We created 100 mock data samples by randomly choosing a subset of events from each of our Monte Carlo samples. From fits to these samples we found that our procedure reproduces the branching fractions without bias, and that the scatter of central values agrees with the uncertainties reported by the fit to better than 15%. These studies were done with the \( |\Delta Q| = 1 \) data included in the vector modes as well as in the pseudoscalar modes. The distribution of likelihoods that we obtained is shown in Fig. 2. For comparison, the likelihood obtained from a comparable fit to the data is also shown. As discussed above, this fits has 406–22 = 384 degrees of freedom. The result from the fit to data is reasonable.

For the actual nominal fit to the data (no \( |\Delta Q| = 1 \) data in the vector modes), we obtained a value \(-2 \ln L = 240.3 \) for 259–19 degrees of freedom. Most bins in the data fit have sizable statistics, so interpretation of \(-2 \ln L \) as a \( \chi^2 \) is reasonable. The probability of \( \chi^2 \) for the fit to the data is 0.48.

In Figs. 3–6 we show the \( M_{m(t)\nu} (\Delta E) \) distributions in the \( \Delta E \) (\( M_{m(t)\nu} \)) signal band for the individual \( q^2 \) regions examined for \( \pi \ell \nu \) and for \( \rho \ell \nu \). For \( \rho \ell \nu \), we show both the distributions with the nominal 1.5 GeV/c minimum lepton momentum requirement and with the more restrictive 2.0 GeV/c requirement of the original CLEO analysis. The fits describe the data in these regions well. The distributions summed over \( q^2 \) for the \( \pi \) and \( \rho \) modes and for \( \omega \ell \nu \) and \( \eta \ell \nu \) are shown in Fig. 7. The \( \omega \ell \nu \) mode remains consistent both with the level expected given the \( \rho \ell \nu \) rate and with pure background. Unless otherwise specified, the normalizations in all figures derive from the fit with the requirement \( p_\ell > 1.5 \text{ GeV/c} \) in the vector modes.

The lepton momentum spectra and \( \cos \theta_{\ell \nu} \) distributions in the \( (M_{m(t)\nu}, \Delta E) \) signal bin are shown in Figs. 8 and 9. This information is not used in the fit, but shows good agreement with the signals preferred in the fit. The \( \pi \pi \) mass distribution for the combined \( \rho \ell \nu \) modes is shown in Fig. 10.

The branching fractions from the nominal fit are summarized in Table III. The results are remarkably stable as the lepton momentum requirement in the vector modes is varied. The greatest variation is observed in the lowest \( q^2 \) interval in the \( \rho \ell \nu \) modes, which we expected because of the larger role that interference between the form factors plays in that region.

Use of a \( \chi^2 \)-based fitting procedure produced similar results, though we saw clearly that low statistics bins had an undue influence on the results of that fitter. Such sensitivity
Carlo samples to account for the uncertainties in the rate and particle identification. In addition, we reweighted the Monte Carlo samples used in the fit. Using independent studies by CLEO for this and other analyses, our modifications reflected the uncertainties in detector simulation and uncertainty in the decay model of the nonsignal B. The breakdown of “ν simulation” into its component parts is given in Table V (and with lepton momentum cuts for vector modes of 1.75 GeV/c and 2.0 GeV/c, in Tables VI and VII, respectively).

We investigated the systematic uncertainties in “ν simulation” by modifying, for each systematic contribution under consideration, the reconstruction output of all of the Monte Carlo samples used in the fit. Using independent studies by CLEO for this and other analyses, our modifications reflected the uncertainties in charged-particle-finding and photon-finding efficiencies, simulation of false charged particles and photons, charged particle momentum resolution, photon energy resolution, hadronic shower simulation, and charged particle identification. In addition, we reweighted the Monte Carlo samples to account for the uncertainties in the rate and was eliminated with the log likelihood minimization. The increase in $-2 \ln L$ from best fit to $B(B^+ \to \eta \ell^+ \nu)$ = 0 is 10.4, corresponding roughly to a 3.2σ statistical significance.

V. EXPERIMENTAL SYSTEMATICS

Table IV summarizes the contributions to the systematic errors for the nominal analysis. The dominant contribution is from uncertainties in “ν simulation,” which includes inaccuracies in detector simulation and uncertainty in the decay model of the nonsignal B. The breakdown of “ν simulation” into its component parts is given in Table V (and with lepton momentum cuts for vector modes of 1.75 GeV/c and 2.0 GeV/c, in Tables VI and VII, respectively).

For many of the variations in the simulation, we expect a cancellation between the change in the signal yield and the change in the efficiency. For complete cancellation ($\sigma_{\text{yield}} = \sigma_{\text{eff}}$, $\sigma_{\text{BR}} = 0$), the additional term amounts to the addition in quadrature of one-third of the change in the yield and in the efficiency. When no cancella-
tion is expected, the additional term is zero. The values for \( \sigma_{\text{yield}} \) and \( \sigma_{\text{eff}} \) are estimated by examining the changes in the "signal bin."

Note that because of correlations between the three \( q^2 \) intervals in a given mode, the sum of the intervals tends to be less sensitive to the systematic variations than the individual intervals themselves.

Consider now the items in Table IV other than "\( \nu \) simulation." We reweight the Monte Carlo sample to allow variation in the relative rates for \( D \ell \nu, D^* \ell \nu, \) and \((Dn\pi)\ell \nu, \) both for resonant \( Dn\pi \) and nonresonant \( Dn\pi. \) We vary the rates by \( \pm 8\%, \pm 6\%, \pm 30\%, \) and \( \pm 30\%, \) respectively. Note that if we completely eliminate any one of these charmed modes except \( D^* \ell \nu, \) the total branching fractions for \( \pi \) and \( \rho \) remain stable within 4\% of themselves, which demonstrates that we are quite insensitive to the details of the poorly measured nonresonant and resonant \((Dn\pi)\) modes. Zeroing \( D^* \ell \nu \) completely causes changes of only 15\%, further demonstrating our insensitivity to the detailed modeling of the \( B \to c \ell \nu \) process.

For the \( B \to X_c \ell \nu \) background, we evaluate two contributions to the systematic uncertainty. First, we vary the nonperturbative parameters of the inclusive spectrum used to drive the \( X_c \ell \nu \) simulation within the uncertainties obtained from the \( B \to X_c \gamma \) analysis that were used in the recent end-point measurement [56,57]. That analysis provides an error ellipse for the HQET parameters \( \lambda_1 \) versus \( \bar{\lambda}, \) and we choose the points on that ellipse that make the maximal change. The second contribution regards uncertainty in the hadronization of the final state light quarks. We change from our model that marries the ISGW II exclusive and OPE inclusive calculations (see the preceding section) to a purely "nonresonant" hadronization procedure (similar to that of JETSET [60]). The hadronization is nonresonant in the sense that single hadron final states (e.g., \( a_1 \ell \nu \)) are not produced. Resonances can appear in the multihadron final state (e.g., \( \rho \pi \ell \nu \)). To avoid overlap of the nonresonant sample with the signal modes, we eliminate \( B \to X_c \ell \nu \) events with a low mass \( \pi \pi \) final state. The uncertainties presented correspond to a minimum \( M_{\pi\pi} \) of 1 GeV. Variation of that threshold over the 0.9–1.1 GeV range results in similar systematic estimates. As a cross-check, we have also used the strictly resonant description of ISGW II, which yields results consistent with our uncertainty estimates.

We have used different normalization schemes for the \( B \to X_c \ell \nu \) background to check the sensitivity of the results under the normalization procedure. If we drop the end-point branching fraction constraint but still allow the normalization to float, we see only minor shifts in the results and the end-point branching fraction predicted by the fit is within one standard deviation of the measured value. We have also used an iterative procedure, where we fix the \( B \to X_c \ell \nu \) normalization in the fit, but update that normalization until the fit’s predicted end-point branching fraction converges to the cen-
down contributes little to the systematic error on $p$ of the small $h$ from $r_{1.43}$, as suggested by Diaz-Cruz @.

Our nominal fit assumed equal production of charged and neutral $B$ mesons: $f_{+}/f_{00}=1$. We varied this fraction over the one standard deviation range indicated by the recent CLEO result $f_{+}/f_{00}=1.04\pm0.08$ [61]. The relationship enters both in the fit to implement the isospin constraint and in the branching fraction calculation to calculate the number of $B^0$ mesons. We used the measured ratio of $B$ meson lifetimes, $\tau_{B^+}/\tau_{B^0}=1.083\pm0.017$, which we varied by one standard deviation to assess the associated uncertainty. The ratio comes into the normalization of the neutral modes versus the charged modes. We have also varied the isospin assumption. In the nominal fit we used a ratio of 2. For the systematic estimate we lowered the $\rho^+:$ $\rho^0$ ratio down to 1.43, as suggested by Diaz-Cruz [62]. The deviation arises from $\rho^0$-$\omega$ mixing coupled with the large $\rho^0$ width. Because of the small $\eta$ and $\omega$ widths, we expect negligible deviation from the ideal factor of 2 for the other two ratios used.

As Table IV shows, uncertainty in the $B\rightarrow X_u \ell \nu$ feed down contributes little to the systematic error on $\pi\ell\nu$ and $\eta\ell\nu$. For the $p\ell\nu$ rate, however, the contribution is substantial.

Our nominal fit assumed equal production of charged and neutral $B$ mesons: $f_{+}/f_{00}=1$. We varied this fraction over the one standard deviation range indicated by the recent CLEO result $f_{+}/f_{00}=1.04\pm0.08$ [61]. The relationship enters both in the fit to implement the isospin constraint and in the branching fraction calculation to calculate the number of $B^0$ mesons. We used the measured ratio of $B$ meson lifetimes, $\tau_{B^+}/\tau_{B^0}=1.083\pm0.017$, which we varied by one standard deviation to assess the associated uncertainty. The ratio comes into the normalization of the neutral modes versus the charged modes. We have also varied the isospin assumption. In the nominal fit we used a ratio of 2. For the systematic estimate we lowered the $\rho^+:$ $\rho^0$ ratio down to 1.43, as suggested by Diaz-Cruz [62]. The deviation arises from $\rho^0$-$\omega$ mixing coupled with the large $\rho^0$ width. Because of the small $\eta$ and $\omega$ widths, we expect negligible deviation from the ideal factor of 2 for the other two ratios used.

The uncertainties related to lepton identification are estimated by varying the measured hadronic fake rates within their uncertainties and by applying the uncertainty in the measurement of the average lepton identification efficiency. Lepton-fake uncertainties are measured in the data as a function of momentum using cleanly tagged hadronic samples, including $K\rightarrow\pi^+\pi^-$ and $D^{*\pm}\rightarrow\pi^+D^0$, $D^0\rightarrow K^\pm\pi^\mp$.

Finally, we assessed our smoothing technique for the continuum data sample. Recall that we use the off-resonance data distribution with relaxed continuum suppression combined with the expected shape change over the fitted $\Delta E$ and $M_{m\ell\nu}$ region that is induced by the relaxation. First of all, fitting without smoothing the off-resonance sample yields results consistent with our nominal fit when the biases expected from downward fluctuations in the off-resonance samples are taken into account [66]. If the few problematic bins are removed from the data, fit results with and without smoothing are essentially identical. The smoothing procedure was introduced to allow the fit to remain unbiased in the presence of these bins. As mentioned above, the predicted shape changes were all consistent with flat (no shape change).

The “MC” prediction used in examining the bias consisted of three components. The primary component was the continuum $q\bar{q}$ component that used JETSET to obtain the initial list of resonances and particles produced in the $e^+e^-$ annihilation. Resonances and short-lived particles were decayed via CLEO’s decay model (QQ) that is based largely on measured branching fractions. The second component arose from continuum events with a hadron misidentified as a lepton. This component was evaluated using data and measured
fake rates. The third component, from $\tau^+\tau^-$, was much smaller [67].

When compared to off-resonance data, our prediction showed excellent agreement in shape for the distributions in ($\Delta E$, $M_{m(e)}$), both with and without the continuum suppression requirements. The largest deviation in the absolute rate prediction was about 20%. Most rate predictions were within 10%. We therefore have good confidence in our ability to predict the ratio of the distributions with and without the suppression. Furthermore, we expect variations in the model parameters—fragmentation function, charm meson decay model, fake rates—that would be consistent with other studies of the continuum to lead to changes that are relatively small compared to the statistical uncertainties. To be somewhat conservative, though, we have changed coherently the ratios of all 45 ($\Delta E$, $M_{m(e)}$) distributions used in the fits, even though the uncertainties are of a statistical nature.

In the $\rho$ modes, there is an additional uncertainty from the unknown contribution of nonresonant $\pi\pi\ell\nu$ decays. While little is known about these decays, we can provide a framework for limiting those contributions through the study of reconstructed $\pi^0\pi^\pm\ell\nu$ decays and the consideration of Bose symmetry, isospin, and angular momentum. The $B\rightarrow X_u\ell\nu$ decay results, before hadronization, in two final-state light quarks. These can have either isospin $I=0$ or $I=1$. Because final-state interactions preserve isospin, a final $\pi\pi$ state is also restricted to $I=0$ or $I=1$. From Bose symmetry considerations, the $\pi\pi$ state must have angular momentum $L$ even for $I=0$ and $L$ odd for $I=1$. Isospin considerations then imply

$I=0,L$ even $\pi^\pm\pi^0: \pi^+\pi^-: \pi^0\pi^0=2:1:0$

$I=1,L$ odd $\pi^+\pi^-: \pi^0\pi^-: \pi^0\pi^0=2:0:1$.

Assuming that the $L=3,5,\ldots$ configurations are suppressed relative to the $L=1$ configuration, we can use $e^+e^-$ scattering data and $\tau$ decay data to conclude that the $I=1,L$ odd component is completely dominated by the $\rho$. A significant nonresonant contribution would therefore come via the $I=0,L$ even channel. With the $I=0$ rate parametrized by $\alpha$, we expect partial widths in the ratios

$\pi^\pm\pi^0: \pi^+\pi^-: \pi^0\pi^0=2:1+2\alpha:\alpha$.

To estimate the systematic due to an unknown nonresonant $\pi\pi\ell\nu$ contribution, we look for a component, after event selection, that could mimic a $\rho\ell\nu$. To constrain such a contribution, we add the mode $\pi^0\pi^0\ell\nu$ to the fit. Procedurally, we generate $\pi^0\pi^0\ell\nu$ using the $\rho$ line shape and the $\rho\ell\nu$ form factors. We then perform fits with the usual isospin constraint on the partial widths ($\rho^\pm:\rho^0=2:1$) replaced with the $\pi\pi$ ratios given above. While the most relevant fit for the extraction of a systematic uncertainty number has the parameter $\alpha$ floating, we also fix $\alpha=0$ to test the fit quality under the assumption that observed $\pi^0\pi^0\ell\nu$ yields are consistent with cross-feed from other modes and the other standard backgrounds.

In the fits, the $\pi^0\pi^0\ell\nu$ mode is treated like the $\omega$ mode. Only the sum of the three $q^2$ intervals contributes to the likelihood, but the signal Monte Carlo simulation is scaled in each $q^2$ interval separately to maintain the above $\pi\pi$ ratios from one interval to the next. Figure 11 shows the projection onto the $m_{\pi^0\pi^0}$ distribution for fits with and without a $\pi^0\pi^0\ell\nu$ signal component. Note that the fit included data only from the three bins in the range $0.485\leq m_{\pi^0\pi^0}$
In addition to the variations above, we have performed numerous systematic checks, including variation of the selection criteria and investigation of electron and muon samples separately. We have also investigated tighter momentum requirements in the pseudoscalar modes. The observed variations were in general consistent within the uncertainties resulting from the statistical changes.

VI. DEPENDENCE OF BRANCHING FRACTIONS ON FORM FACTORS

In the original measurement of the exclusive charmless branching fractions [3], there were two roughly comparable contributions to the branching fraction errors from the form-factor uncertainties. The first contribution resulted because the efficiency varied as a function of \( q^2 \) (inescapable with a lepton momentum cut), and the data were lumped into a single \( q^2 \) bin. Because we now extract the rates independently in three separate \( q^2 \) ranges, this analysis should see a significant reduction in this effect. The second contribution resulted because there was significant \( q^2 \) dependence to the cross-feed rates between the pseudoscalar and the vector modes. Again, since we extract the rates independently as a function of \( q^2 \), this dependence should be reduced.

We have estimated the model dependence based on changes of the branching fractions under variation of the form-factor calculation. The previous analysis [3] found that the error on the branching fraction obtained from comparison of models was larger than that obtained by variation of a particular form-factor parametrization within the published uncertainties (when given). Tables IX and X show the variation in \( B(B^0 \rightarrow \pi^- \ell^+ \nu) \) and \( B(B^0 \rightarrow \rho^- \ell^+ \nu) \), respectively, as the \( \pi^- \) and vector form factors are varied. We have included in the set of models those which have the most ex-
TABLE IV. Contributions to the systematic error (%) in each total and partial branching fraction (B). Simulation of the detector and the second B contribute to υ simulation.

<table>
<thead>
<tr>
<th>Systematic</th>
<th>$B_{total}$</th>
<th>$q^2$ interval (GeV^2)</th>
<th>$\geq$16</th>
<th>$B_{total}$</th>
<th>$q^2$ interval (GeV^2)</th>
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<th>$\eta$</th>
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<tbody>
<tr>
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<td></td>
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<td></td>
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<tr>
<td>$B \rightarrow D(D^*+/D^{**+}/D^{NR}X)$</td>
<td>1.7</td>
<td>2.5</td>
<td>1.9</td>
<td>3.2</td>
<td>2.0</td>
<td>21.4</td>
<td>4.7</td>
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<td>$B \rightarrow X_{s} \ell\nu$ feed down</td>
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<td>2.0</td>
<td>3.0</td>
<td>10.0</td>
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<tr>
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<tr>
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</table>

TABLE V. Percentage change in results for a fit with a modified simulation relative to a fit to the nominal MC simulation for each of the variations contributing to the simulation systematic uncertainty. The vector modes were analyzed with the requirement $p_{\ell}>1.5$ GeV/c. The last row shows the quadrature sum of the changes.

<table>
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<th>Variation</th>
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the variations contributing to the simulation systematic uncertainty. The vector modes were analyzed with the requirement $p_{\ell} > 1.75 \text{ GeV}/c$. The last row shows the quadrature sum of the changes.

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The $\rho \ell \nu$ variation remains significant, though again this analysis shows essentially no dependence on the $\pi \ell \nu$ form factor. The overall uncertainty of the form factors has reduced to about 80% of the original CLEO $\rho \ell \nu$ measurement [3] (which had a smaller form-factor dependence than the 2000 CLEO $\rho \ell \nu$ analysis [4]). As one tightens the lepton momentum requirement, the model dependence increases slightly over the range we have studied. As expected, the lowest $q^2$ interval shows the greatest sensitivity (fractionally) to the variation in the range. For a given model, the variation of the total branching fraction as the lepton momentum requirement is varied is small compared to the variation among models for a given momentum requirement. (The rms variation of the former is about 30% of the rms variation of the latter.) We speculate that the dominant model dependence likely arises from our $\cos \theta_{W}\ell > 0$ requirement, which we applied to suppress $b \rightarrow c$ background. Either finer $q^2$ binning or an alternate means of background suppression would provide a route for further reduction of the form-

<table>
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<th>$q^2 &lt; 8$</th>
<th>$8 \leq q^2 &lt; 16$</th>
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factor dependence.

For the $\eta \ell \nu$ branching fraction, we find a dependence of $0.04 \times 10^{-4}$ from variation of the $\pi \ell \nu$ form factors and $0.01 \times 10^{-4}$ from variation of the $\rho \ell \nu$ form factors. The only $\eta \ell \nu$ form factor that we consider is ISGW II [31]. However, the $\eta \ell \nu$ analysis presented here is almost identical to the original $\pi \ell \nu$ analysis. We therefore take the form-factor dependence of 10% found in that analysis as an estimate of the uncertainty from the $\eta$ form factors. As the $\rho \ell \nu$ form factors contributed substantially to the 10% uncertainty in the previous analysis, yet contribute negligibly to $\eta \ell \nu$, the 10% should be a conservative estimate.

The results presented here agree well with the previous CLEO measurements and the recent BABAR $\rho \ell \nu$ measurement. The results of the original CLEO measurement [3] are superseded by this measurement. The results of the CLEO 2000 measurement [4] are essentially statistically independent of those presented here.

VII. EXTRACTION OF $|V_{ub}|$ AND DISCRIMINATION OF MODELS

We extract $|V_{ub}|$ from the measured rates for $\pi \ell \nu$ only, for $\rho \ell \nu$ only, and then by using the combined information from those two modes. In all cases, the $|V_{ub}|$ extraction is based on the results from the analysis requiring $p_t > 1.5$ GeV/c in the vector modes. We use a $B^0$ lifetime of $(1.542 \pm 0.016)$ ps [58].

A. $|V_{ub}|$ from $B \to \pi \ell \nu$

For $\pi \ell \nu$, we first explore fitting $q^2$ distributions from various form-factor predictions to the measured rates in the three $q^2$ bins. To be self-consistent, we extract $|V_{ub}|$ for a particular form factor using the rates from the fit with that model. In practice, as we have seen, this makes little difference in the $\pi$ modes in this analysis. Since each model predicts the total rate modulo $|V_{ub}|$, $|V_{ub}|$ becomes the one free parameter for the fit that normalizes the prediction to the observed rates. The quality of the fit measures how well the form-factor shape describes the data, so it provides one means of discrimination among form factors. The results of this procedure are summarized in Table XI. For the three calculations that have been used for both efficiency and $|V_{ub}|$ extraction, the data rates with the best fits for each predicted form factor are shown in Fig. 12. The probability of $\chi^2$ in our various fits for the ISGW II model varies between 1% and 3%, indicating that this model is likely to be less reliable for determination of $|V_{ub}|$ from $\pi \ell \nu$. Note further that the spread among the central values from the various calculations is fairly small relative to the uncertainties quoted in the calculations themselves.

Because the extracted rates in the $q^2$ intervals are now essentially independent of the $\pi \ell \nu$ form factor, one can extract $|V_{ub}|$ from our results for form factors not considered here. We provide in Appendix B a detailed methodology for doing so.

To determine the effect of the systematic uncertainties, we repeat the above fit using the three $q^2$ rates obtained from the branching ratio fit after each systematic variation. This procedure automatically accounts for correlations among the three intervals. We then increase the uncertainty for each variation by one-half of the fractional error introduced by the second term in Eq. (7). The factor of one-half arises from the square root involved in extraction of $|V_{ub}|$ from the rate.

As we discuss below, each of the form-factor calculations used to extract $|V_{ub}|$ from the full $q^2$ range has some measure of model dependence. We determine a systematic error in $|V_{ub}|$ from the quoted theoretical uncertainty in form

![Graph](image-url)
Due to the poor agreement of the ISGW II form-factor normalizations, with the following procedure. For each form factor used, we recalculate $|V_{ub}|$ when we increase or decrease the form-factor normalization by one standard deviation. Due to the poor agreement of the ISGW II form factor with the $\pi\ell\nu$ data in conjunction with the some-

what ad hoc assumptions about the form-factor $q^2$ dependence in that mode, we drop ISGW II from consideration. From the others, we find the minimum value $V_{\min}$ and the maximum value $V_{\max}$. We then assign an asymmetric error of 70% of the deviation relative to the nominal central value.

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<td>UKQCD'98</td>
<td>2.23±0.32</td>
<td>0.45±0.18</td>
<td>1.28±0.25</td>
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<td>240.4</td>
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<tr>
<td>1.7×RMS$_{\pi,FF}$</td>
<td>0.44</td>
<td>0.11</td>
<td>0.24</td>
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<table>
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<th>$p$ model</th>
<th>$B_{total}$</th>
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<th>$B_{&lt;16}$</th>
<th>$B_{&gt;16}$</th>
<th>$-2 \ln L$</th>
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<tbody>
<tr>
<td>Ball'01</td>
<td>Ball'98</td>
<td>2.29±0.35</td>
<td>0.62±0.22</td>
<td>1.11±0.25</td>
<td>0.56±0.10</td>
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<tr>
<td>Ball'01</td>
<td>ISGW2</td>
<td>1.89±0.27</td>
<td>0.38±0.13</td>
<td>0.98±0.22</td>
<td>0.54±0.09</td>
<td>243.4</td>
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<tr>
<td>Ball'01</td>
<td>Melikhov'00</td>
<td>2.66±0.38</td>
<td>0.48±0.17</td>
<td>1.36±0.31</td>
<td>0.83±0.14</td>
<td>244.6</td>
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<td>UKQCD'98</td>
<td>2.15±0.32</td>
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<td>1.07±0.24</td>
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<td>0.15</td>
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<td>0.21</td>
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</table>
TABLE XI. $|V_{ub}|$ extracted from fits to the rates measured in the three $q^2$ intervals for a variety of form factors for $\pi^0\ell\nu$. The table indicates form-factor calculation, $|V_{ub}|$ with statistical error only, predicted $\Gamma_{\pi^0}/|V_{ub}|^2$ with the estimated theoretical uncertainty, the $\chi^2$ for the fit, and the probability of $\chi^2$ given the two degrees of freedom.

| $\pi$ model | $\rho$ model | $|V_{ub}| \times 10^3$ | $\Gamma_{\pi^0}/|V_{ub}|^2$ (ps$^{-1}$) | Fit $\chi^2$ | $P(\chi^2)$ |
|-------------|--------------|----------------|---------------------------------|-------------|-------------|
| Ball’01     | Ball’98     | 3.21 ± 0.21  | 8.4$^{+3.5}_{-2.4}$             | 1.0         | 0.61        |
| KRWWY$^a$ [22] | Ball’98     | 3.40 ± 0.23  | 7.3 ± 2.5                       | 5.3         | 0.07        |
| ISGW2       | Ball’98     | 2.90 ± 0.20  | 9.6 ± 4.8                       | 7.3         | 0.03        |
| SPD         | Ball’98     | 2.96 ± 0.19  | 9.6 ± 2.9                       | 4.0         | 0.14        |
| Ball’01     | Ball’98     | 3.21 ± 0.21  | 8.4$^{+3.5}_{-2.4}$             | 1.0         | 0.61        |
| Ball’01     | ISGW2       | 3.31 ± 0.20  | 8.4$^{+3.5}_{-2.4}$             | 1.2         | 0.55        |
| Ball’01     | Melikhov’00 | 3.18 ± 0.21  | 8.4$^{+3.5}_{-2.4}$             | 0.9         | 0.63        |
| Ball’01     | UKQCD’98    | 3.24 ± 0.20  | 8.4$^{+3.5}_{-2.4}$             | 1.1         | 0.59        |

$^a$ Uses rates determined with the Ball’01 form factor.

Again for direct comparison with other experiments, taking one-half, rather than 70%, as the scale factor for estimating the uncertainties yields $|V_{ub}| = (3.21 ± 0.21 ± 0.14 ± 0.07) \times 10^{-3}$.

Note that the error on $|V_{ub}|$ from the uncertainty in the rates under variation of form factors is completely dwarfed by the error arising from uncertainty in the theoretical normalization of the form factor.

Our second, preferred, method for determining $|V_{ub}|$ attempts to reduce the number of modeling assumptions and hence to provide a more robust uncertainty estimate. We therefore limit our consideration to form factors determined from LCSR and from LQCD calculations, which are QCD-based approaches. These calculations, however, are each only valid over a restricted $q^2$ region. The QCSR calculations are expected to break down for $q^2 \geq 16$ GeV$^2$, while the current LQCD calculations are valid only for $q^2 \geq 16$ GeV$^2$. Extrapolation outside of these ranges therefore introduces a dependence on the form used for the extrapolation. This introduces another uncertainty that is difficult to assess. To minimize this uncertainty, we extract $|V_{ub}|$ from these more restricted regions. For LQCD, we determine $|V_{ub}|$ from the measured rate and the calculated rate in the range $q^2 \geq 16$ GeV$^2$. For LCSR, we determine $|V_{ub}|$ by fitting the calculated LCSR rates to the measured rates in the two $q^2$ intervals below 16 GeV$^2$. The results are shown in Table XII.

To produce a final LQCD result for the $q^2 \geq 16$ GeV$^2$ region, we take a statistically weighted average of the different LQCD results. To the precision quoted, we obtained identical results if we based the statistical weights on the upper, the lower, or the average of the asymmetric statistical errors quoted in Table XII. We assume the systematic errors are completely correlated among the different calculations: if $\sigma_i$ is the statistical weight used in the average for calculation $i$ and $\hat{\sigma}_i$ is the fractional systematic error for that calculation, then the total fractional systematic error $\hat{\sigma}$ assigned to the average is $\hat{\sigma} = \Sigma \alpha_i \hat{\sigma}_i$. The theoretical errors quoted in Table XII do not include any uncertainty from the quenched approximation, which is estimated to be in the 10% to 20% range. We add an additional 15% in quadrature to the sys-
tect the average with a weight of the procedure laid out in Appendix C. The LQCD value en-
gaged, according to experimental systematics taken into account, according to
gives similar theoretical uncertainties. Using the fractional errors from the LCSR calculations alone
We take this as the more reliable determination of $|V_{ub}|$ from our complete data in this mode.

The variations in $|V_{ub}|$ and our averages are illustrated in Fig. 13.

![FIG. 13. Values for $|V_{ub}|$ obtained from $\pi \ell \nu$ using the entire $q^2$ range for the various form-factor calculations (top block), using LQCD for $q^2 \geq 16$ GeV$^2$ (second block), using LCSR for $q^2 < 16$ GeV$^2$ (third block), and using our average of the last two (bottom block) for $\pi \ell \nu$ only and for $\pi \ell \nu$ and $\rho \ell \nu$ combined. In all cases, the top bar indicates the statistical and all the experimental systematics (combined in quadrature), and the lower bar indicates the approximate “one standard deviation” range of motion due to the theoretical uncertainties.](072003-18)
Estimating the uncertainties as described in the preceding Table XIII, we take the Ball’98 results as the central value. For an estimate of $|V_{ub}|$ based on the models and fits in Table XIII, we take the Ball’98 results as the central value. Estimating the uncertainties as described in the preceding section, we obtain

$$|V_{ub}| = (2.90 \pm 0.21^{+0.31}_{-0.36} \pm 0.73) \times 10^{-3},$$

where the errors are statistical, experimental systematic, and the estimated uncertainties from 70% of the total spread in the results as we vary the $\rho \ell \nu$ form-factor calculations over ±1 standard deviation, respectively. This estimate is similar to, though somewhat larger than, that obtained from the quoted Ball’98 uncertainty.

Restricting ourselves to the theoretically more reliable use of LQCD for $q^2 \geq 16$ GeV$^2$ and LCSR for $q^2 < 16$ GeV$^2$, we have only the two results listed in Table XIV. In addition to the theoretical uncertainty quoted for UKQCD’98, we add an additional 20% in quadrature as an estimate of the quenching uncertainty. This is larger than for the $\pi \ell \nu$ case both because the $\rho$ is a broad resonance and because of the potential for larger biases from quenching given the interference between the various form factors. We also apply our reinterpretation of symmetric theoretical errors on the rate as symmetric errors on the amplitude. The results in the two $q^2$ intervals are thus

$$|V_{ub}|_{q^2 \geq 16 \text{ GeV}^2} = (3.34 \pm 0.32^{+0.27}_{-0.36} + 0.50) \times 10^{-3}$$

and

$$|V_{ub}|_{q^2 < 16 \text{ GeV}^2} = (2.67 \pm 0.27 + 0.38^{+0.47}_{-0.42} + 0.35) \times 10^{-3}.$$  

We average the LQCD and LCSR results, with correlated experimental systematics taken into account. We again employ the procedure described in Appendix C. The optimal weight for combining the two intervals, treating the systematic uncertainties as completely correlated, is $\alpha_{\rho} = 0.5$. We find

$$|V_{ub}| = (3.00 \pm 0.21^{+0.29}_{-0.35} + 0.49 + 0.28) \times 10^{-3}.$$

The errors are statistical, experimental systematic, theoretical systematic based on the LQCD and LCSR uncertainties, and $\rho \ell \nu$ form-factor shape uncertainty. To be conservative, we have assigned the latter error based on the variation seen in the total branching fraction in this mode. The contribution

| TABLE XIV. Values for $|V_{ub}|$ obtained using form factors (FF) from light-cone sum rules in the $q^2$ interval 0–16 GeV$^2$ (first row) and from LQCD for $q^2 \geq 16$ GeV$^2$ (second row). Only the statistical errors are indicated. The data rates obtained using Ball’01 for $\pi \ell \nu$ and Ball’98 for $\rho \ell \nu$ were used as the input for all values obtained. |
|---|---|---|---|---|
| $\rho$ | FF | $|V_{ub}| \times 10^3$ | $\Gamma_{p\ell\nu}^\text{th}/|V_{ub}|^2$ (ps$^{-1}$) | Fit $\chi^2$ | $P(\chi^2)$ |
| Ball’98 | | 2.67±0.27 | 14.2±4.3 | 4.5 | 0.03 |
| UKQCD’98 | | 3.34±0.32 | 2.9±0.62 | 0.17 | 0.07 |

| TABLE XIII. $|V_{ub}|$ extracted from fits to the rates measured in the three $q^2$ intervals for a variety of form factors for $\rho \ell \nu$. The table indicates form-factor calculation, $|V_{ub}|$ with statistical error only, predicted $\Gamma_{p\ell\nu}^\text{th}/|V_{ub}|^2$ with the estimated theoretical uncertainty, the $\chi^2$ for the fit, and the probability of $\chi^2$ given the two degrees of freedom. |
|---|---|---|---|---|
| $q^2$ (GeV$^2$) | $|V_{ub}| \times 10^3$ | $\Gamma_{p\ell\nu}^\text{th}/|V_{ub}|^2$ (ps$^{-1}$) | Fit $\chi^2$ | $P(\chi^2)$ |
| Ball’01 | Ball’98 | 2.90±0.21 | 16.9±5.1 | 7.6 | 0.02 |
| Ball’01 | UKQCD’98 | 2.88±0.20 | 16.5±3.3 | 5.2 | 0.08 |
| ISGW2 | Ball’98 | 2.90±0.21 | 16.9±5.1 | 7.6 | 0.02 |
| SPD | Ball’98 | 2.90±0.21 | 16.9±5.1 | 7.6 | 0.02 |
FIG. 15. Values for $|V_{ub}|$ obtained from $\rho \ell \nu$ using the entire $q^2$ range for the various form-factor calculations (top block), using LQCD for $q^2 > 16$ GeV$^2$ (second block), using LCSR for $q^2 < 16$ GeV$^2$ (third block), and our average of the last two (bottom block) for $\rho \ell \nu$ only and for $\pi \ell \nu$ and $\rho \ell \nu$ combined. In all cases, the top bar indicates the statistical and all the experimental systematics (combined in quadrature), and the lower bar indicates the approximate “one standard deviation” range of motion due to the theoretical uncertainties.

from the $\pi \ell \nu$ form-factor shape is negligible. Again, we take this as our preferred method of extracting $|V_{ub}|$ from our $\rho \ell \nu$ data.

The $|V_{ub}|$ results obtained from $\rho \ell \nu$ are shown in Fig. 15.

C. $|V_{ub}|$ from a combination of $B \to \pi \ell \nu$ and $B \to \rho \ell \nu$

We have averaged the $|V_{ub}|$ determinations obtained separately from the $B \to \pi \ell \nu$ and $B \to \rho \ell \nu$ modes. For this average, we considered only the results obtained using the LCSR and LQCD calculations applied to the $q^2 > 16$ GeV$^2$ and $q^2 \geq 16$ GeV$^2$ results, respectively. The averaging procedure amounts to the determination of the optimal weight $\beta$ to be applied to the LCSR and LQCD average obtained from $B \to \pi \ell \nu$ relative to that obtained from $B \to \rho \ell \nu$ (see Appendix C). We held the values $\alpha_\pi$ and $\alpha_\rho$, each of which determines the weight of the LQCD result relative to the LCSR result in the individual mode, fixed at the optimal values found in the preceding sections. The weight $\beta = 0.7$ provided the optimal combination. With this weighting, we find

$$|V_{ub}| = (3.17 \pm 0.17 \pm 0.16 + 0.53 \pm 0.03) \times 10^{-3}. \quad (16)$$

The errors are statistical, experimental systematic, theoretical systematic based on the LQCD and LCSR uncertainties, and $\rho \ell \nu$ form-factor shape uncertainty, respectively. Note that because of cross-feed among the modes considered, the $\pi \ell \nu$ and $\rho \ell \nu$ modes are anticorrelated, resulting, in particular, in the minimal dependence of the average result on the $\rho \ell \nu$ form-factor shape.

VIII. SUMMARY

With a sample of $9.7 \times 10^8 B \bar{B}$ pairs, we have studied $B$ decays to $\pi \ell \nu$, $\rho \ell \nu$, $\omega \ell \nu$, and $\eta \ell \nu$, where $\ell = e$ or $\mu$. From the combination of a broad momentum range for the charged lepton momentum and independent extraction of rates in three separate $q^2$ intervals, we were able to reduce the uncertainties from modeling within the form-factor calculations. For the decay $B^0 \to \pi^- \ell^+ \nu$, we have determined the branching fractions

$$B(0 \leq q^2 < 8 \ {	ext{GeV}}^2) = (0.43 \pm 0.11 \pm 0.05 \pm 0.004 \pm 0.01) \times 10^{-4}$$

$$B(8 \leq q^2 < 16 \ \text{GeV}^2) = (0.65 \pm 0.11 \pm 0.07 \pm 0.01 \pm 0.03) \times 10^{-4}$$

$$B(q^2 \geq 16 \ \text{GeV}^2) = (0.25 \pm 0.09 \pm 0.04 \pm 0.01 \pm 0.03) \times 10^{-4}. \quad (17)$$

Combining these rates and taking into account correlated systematic uncertainties, we obtain

$$B(B^0 \to \pi^- \ell^+ \nu) = (1.33 \pm 0.18 \pm 0.11 \pm 0.01 \pm 0.07) \times 10^{-4}. \quad (18)$$

where the errors are statistical, experimental systematic, the estimated uncertainties from the $\pi \ell \nu$ form factor, and those from the $\rho \ell \nu$ form factors, respectively.

For the decay $B^0 \to \rho^- \ell^+ \nu$, we have determined the branching fractions

$$B(0 \leq q^2 < 8 \ \text{GeV}^2) = (0.43 \pm 0.20 \pm 0.23 \pm 0.09 \pm 0.01) \times 10^{-4}$$

$$B(8 \leq q^2 < 16 \ \text{GeV}^2) = (1.24 \pm 0.26 + 0.07 \pm 0.03 \pm 0.22 \pm 0.004) \times 10^{-4}$$

$$B(q^2 \geq 16 \ \text{GeV}^2) = (0.50 \pm 0.10 + 0.08 \pm 0.11 \pm 0.19 \pm 0.004) \times 10^{-4}. \quad (19)$$

Combining these rates, again taking into account correlated systematic uncertainties, we obtain

$$B(B^0 \to \rho^- \ell^+ \nu) = (2.17 \pm 0.34 + 0.47 \pm 0.54 \pm 0.41 \pm 0.01) \times 10^{-4}. \quad (20)$$

where the errors are statistical, experimental systematic, the estimated uncertainties from the $\rho \ell \nu$ form factors, and those from the $\pi \ell \nu$ form factor, respectively.

When the theoretical uncertainties that result from form-factor $q^2$ dependence are evaluated in a common fashion, the branching fractions obtained in this analysis have uncertainties from the form-factor $q^2$ dependence that are reduced by about a factor of 2 compared to previous $\rho \ell \nu$ analyses [3, 463]. These uncertainties are almost eliminated for the $\pi \ell \nu$ branching fraction.
We see evidence for the decay $B^+ \rightarrow \eta \ell^+ \nu$ with a statistical significance corresponding roughly to $3.2\sigma$. The rate we obtain,

$$\mathcal{B}(B^+ \rightarrow \eta \ell^+ \nu) = (0.84 \pm 0.31 \pm 0.16 \pm 0.09) \times 10^{-4},$$

(21)

is consistent, within sizable errors, with that expected from the measured pion rate and isospin relations. Only an ISGW II form factor has been examined, and a 10% model dependence uncertainty has been assigned based on the previous CLEO $\pi \ell \nu$ analysis. The final error quoted combines this estimate with the dependence on the $\pi \ell \nu$ and $\rho \ell \nu$ form factors.

From the $\pi \ell \nu$ $q^2$ behavior that we have observed, we find the ISGW II form factor for $\pi \ell \nu$ consistent with data at only the 3% level.

By fitting LQCD and LCSR calculations to the observed $q^2$ behavior in $\pi \ell \nu$, restricting each calculation to its valid $q^2$ range, and then combining the results, we extract

$$|V_{ub}| = (3.24 \pm 0.22 \pm 0.13 \pm 0.39 \pm 0.09) \times 10^{-3},$$

(22)

where the errors are statistical, experimental systematic, the estimated uncertainties from the $\pi \ell \nu$ form-factor shape and normalization, and those from the $\rho \ell \nu$ form factors’ shapes, respectively. From a similar analysis of the $\rho \ell \nu$ mode, we obtain

$$|V_{ub}| = (3.00 \pm 0.21 \pm 0.29 \pm 0.35 \pm 0.28) \times 10^{-3}.$$ 

(23)

The errors are statistical, experimental systematic, theoretical systematic based on the LQCD and LCSR uncertainties, and $\rho \ell \nu$ form-factor shape uncertainty, respectively. In general, the $\rho \ell \nu$ form-factor calculations did not agree as well with the observed $\rho \ell \nu$ data as did the $\pi \ell \nu$ form-factor calculations with the $\pi \ell \nu$ data.

Combining these two modes for an overall result from this analysis, we obtain

$$|V_{ub}| = (3.17 \pm 0.17 \pm 0.16 \pm 0.39 \pm 0.03) \times 10^{-3}.$$ 

(24)

Given the manner with which the theoretical uncertainties have been estimated, the quoted values should be interpreted as being closer in spirit to “one standard deviation” than to “the allowed range.”

These results trade off the potential statistical gain over the previous CLEO analyses in favor of relaxation of theoretical constraints. Had we fixed the relative rate in the three $q^2$ intervals in the $\pi \ell \nu$ and $\rho \ell \nu$ modes, a more pronounced improvement in statistical precision would have resulted. By relaxing the constraint, on the other hand, we have minimized our reliance on modeling in extraction of rates and of $|V_{ub}|$.

These results supersede the $\pi \ell \nu$ and $\rho \ell \nu$ results obtained in Ref. [3]. They agree, within measurement uncertainties, with the CLEO 2000 $\rho \ell \nu$ result [4] and with the recent BABAR $\rho \ell \nu$ analysis [63].

The results for $|V_{ub}|$ obtained here are compatible with the results obtained from the recent CLEO end-point measurement [57]. The estimated theoretical uncertainties remain sizable for both $\pi \ell \nu$ and $\rho \ell \nu$, and there remain uncertainties in the estimates themselves. We therefore do not average these results, but view the compatibility as an indication that the uncertainties have not been appreciably underestimated.

Significant progress in extraction of $|V_{ub}|$ from exclusive decays will require a major improvement in theory.

**ACKNOWLEDGMENTS**

We thank A. Kronfeld, J. Simone, T. Onogi, T. Feldmann, P. Kroll, C. Maynard, and D. Melikhov for assistance with form factors. We gratefully acknowledge the effort of the CESR staff in providing us with excellent luminosity and running conditions. M. Selen thanks the Research Corporation, and A.H. Mahmoud thanks the Texas Advanced Research Program. This work was supported by the National Science Foundation and the U.S. Department of Energy.

**APPENDIX A: DESCRIPTION OF EXPERIMENTAL SYSTEMATIC UNCERTAINTY DETERMINATION**

The techniques employed in this analysis rest fundamentally on complete, accurate reconstruction of all particles from both $B$ decays in an event. As a result, systematic uncertainty estimates that reflect uncertainties in the detector simulation must account for the reliability with which an entire event can be reconstructed, not just the signal particles. For example, if there is a residual uncertainty in the track reconstruction efficiency, the signal efficiency will not only be affected by incorrectly assessing the loss of the signal mode particles, it will also be affected by “misreconstruction” of the neutrino four-momentum. Furthermore, the rate at which background samples can smear into the signal region is also affected by the overall misreconstruction.

We therefore estimate the systematic uncertainties due to detector modeling by modifying each reconstructed Monte Carlo event in each signal and background sample. For each study, the size of the variation has generally been determined by independent comparisons of data and Monte Carlo. The following list describes the variations that enter the systematic determination.

**Tracking efficiency.** We have limited our uncertainty in track-finding efficiency for high (above 250 MeV/c) and low momentum tracks to be under 0.5% and 2.6%, respectively. These limits were obtained with hadronic samples, and therefore include any discrepancies in the interaction cross sections. To determine the systematic error from the uncertainty in tracking efficiency, we apply an additional inefficiency of 0.75% and 2.6% to each high momentum track and to each low momentum track, respectively, in the simulation.

**Tracking resolution.** We increase the mismeasurement of each momentum component for each reconstructed charged particle by 10% of itself, which is outside the range for which core distributions agree, but compensates for discrepancies in the tails.

**$\gamma$ efficiency.** We have limited our uncertainty in photon reconstruction efficiency to 2%. In our studies, we have ac-
tually applied an additional 3% efficiency loss per photon, then scaled the observed shifts back by 2/3.

\textit{γ resolution.} We also degrade the photon energy resolution by 10% of itself.

\textit{Split-off simulation.} Studies of $\gamma\gamma \rightarrow K_S K_S$ have indicated that the combination of mismodeling the physics processes and hadronic showers leads to an excess of isolated reconstructed showers (split offs) at the rate of 0.03/hadron in data relative to the Monte Carlo simulation. To estimate the potential effect on our analysis, we interpret the entire excess as mismodeling of the hadronic showers, and add showers at this rate to each of our Monte Carlo samples.

\textit{Split-off rejection.} We bias our neural net parameter, which is derived from the distribution of energy within the crystals in the shower relative to the primary impact point of a “parent” charged hadron, to move photon-like results in the Monte Carlo simulation towards hadronic-shower-like results. We limit the variations based on data and Monte Carlo comparisons of the parameter as a function of shower energy.

\textit{K_L showers.} In our simulation of $K_L$ showers, we increase the energy deposited in our CsI calorimeter. The variation is based on data and Monte Carlo comparisons of the energy deposited by $K^\pm$ showers after correction for the minimum-ionizing component.

\textit{K_L production.} By comparing the data and Monte Carlo $K_S$ energy spectrum and yield, we found that our $K_L$ rate needed to be decreased by $(7.2 \pm 1.0)\%$, and that no correction was needed for the spectrum. The nominal analysis reweights events with $K_L$ accordingly, and we vary the weight according to its uncertainty to estimate the systematic contribution.

\textit{Extra ν production.} An important source of background is events that contain both a $b \rightarrow c \ell \nu$ decay and a $c \rightarrow s \ell \nu$ decay, where the latter can originate with either $B$ meson in the event. We reweight the Monte Carlo sample so that the lepton momentum spectrum from secondary charm decay agrees with a spectrum obtained by convoluting a recent measurement of the charm meson momentum spectrum from $B$ decay [64] with the MARK III measurement of the inclusive lepton momentum spectrum from charm decay [65]. The nominal result is corrected based on this procedure. To estimate the systematic uncertainty, we define spectrum “envelopes” and reweight our Monte Carlo samples to match this spectrum. The envelopes were defined by throwing 500 toy Monte Carlo spectra in which all experimental inputs were varied according to their uncertainties and finding the variation within each momentum bin that contained 68% of the toy spectra.

\textit{Particle ID.} We simultaneously shift all $dE/dx$ and time-of-flight distributions in the simulation by 1/4 and 1/2 of the intrinsic resolution, respectively. We take the full variation we observe as our uncertainty, even though this procedure leads to a very conservative systematic estimate.

For each of these variations, we modify or reweight each event in each Monte Carlo sample in a full reanalysis of these samples. The set of modified samples for each variation replaces the nominal samples input to the branching fraction fit. For each variation, the shifts in the fit results provide the first input into the systematic estimates on the branching fractions for that variation. We can view the shifts in results as arising from two components: a change in the signal efficiency and a change in the predicted background level. These changes tend to cancel in the total shift: a variation that reduces the signal reconstruction efficiency also simultaneously increases the background level (and reduces the signal yield from the fit). As the main text describes, we increase our systematic estimate to allow for imperfections in the predicted cancellation.

\textbf{APPENDIX B: EXTRACTION OF $|V_{ub}|$ FROM THE MEASURED $d\Gamma(B^0 \rightarrow \pi^- \ell^+ \nu)/dq^2$ DATA WITH FUTURE FORM-FACTOR CALCULATIONS}

The branching fractions in the three $q^2$ ranges for $B \rightarrow \pi \ell \nu$ exhibit very little dependence on the precise form factors used to extract the branching fractions. The results can therefore be reliably used to obtain values for $|V_{ub}|$ using future $B \rightarrow \pi \ell \nu$ form-factor calculations that are improved over those used in this paper. This appendix provides the detail needed to ascertain the proper experimental uncertainties for such an extraction using the same fitting technique presented above. The main difficulty stems from proper evaluation of the experimental uncertainties because of correlations (both positive and negative) among the results for the three ranges. The correlations arise both statistically from the fitting procedure used to extract the three rates and systematically as we vary the details of the simulation.

To extract a central value of $|V_{ub}|$, we perform a $\chi^2$ fit to the nominal branching fractions from the three $q^2$ intervals listed in Table XV. This $|V_{ub}|$ fit includes the correlation coefficients among the rates from the branching fraction fit to the data: $\rho_{12} = -0.035$, $\rho_{13} = 0.003$, and $\rho_{23} = -0.037$.

To evaluate the error arising from simulation uncertainties (“ν simulation” in Table IV) on the results, we redo our $\chi^2$ fit for $|V_{ub}|$ using the new rates listed in Table XV for each variation. For the results presented here, we have used the correlation coefficients from the branching fraction fit to the data for each variation. In practice, the coefficients remain stable enough that using the nominal coefficients in all fits is sufficient. The change relative to the nominal $|V_{ub}|$ result provides the first input to the uncertainty estimate. For the uncertainty estimate in $K_L$ production and secondary ν production, we take the average of the “up” and “down” shifts as our overall estimate. To allow for misestimation of correlated changes between background levels and signal efficiencies in the results (see main text), we increase the fractional uncertainty on $|V_{ub}|$ from each variation by adding in quadrature the quantities listed in Table XVI. Finally, the ν efficiency uncertainty should be scaled back to 2/3 of the value found above. We combine all of the uncertainties in quadrature to arrive at the total “ν simulation” systematic for $|V_{ub}|$.

We evaluate the uncertainty from our modeling of the $B \rightarrow X_S \ell \nu$ backgrounds in much the same fashion. The fit variations that we have used for this purpose are listed in Table XVII. An earlier version of our $B \rightarrow X_S \ell \nu$ generator was used in the study, and the table also shows the “nomi-
TABLE XV. Central values and statistical uncertainties for $B^0 \rightarrow \pi^- \ell^+ \nu$ branching fractions for the nominal fit and for each systematic variation of the Monte Carlo samples input to the fit. The detector-related systematic uncertainties in $|V_{ub}|$ are obtained by fitting the results from the relevant set of $q^2$ intervals for each systematic study. The total branching fraction is shown as well for completeness. All results were obtained using the Ball’01 form factor for the $\pi \ell \nu$ modes and the Ball’98 form factors for the $\rho \ell \nu$ modes.

<table>
<thead>
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<th>Systematic change</th>
<th>Total $10^3 \times B(B^0 \rightarrow \pi^- \ell^+ \nu)$</th>
<th>$0 \leq q^2 &lt; 8$ GeV$^2$</th>
<th>$8 \leq q^2 &lt; 16$ GeV$^2$</th>
<th>$16$ GeV$^2 \leq q^2 &lt; q^2_{\text{max}}$</th>
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<tr>
<td>Nominal</td>
<td>1.327 ± 0.177</td>
<td>0.431 ± 0.106</td>
<td>0.651 ± 0.105</td>
<td>0.245 ± 0.094</td>
</tr>
<tr>
<td>$\gamma$ eff.</td>
<td>1.348 ± 0.194</td>
<td>0.476 ± 0.117</td>
<td>0.674 ± 0.117</td>
<td>0.198 ± 0.103</td>
</tr>
<tr>
<td>$\gamma$ resol.</td>
<td>1.379 ± 0.183</td>
<td>0.445 ± 0.111</td>
<td>0.686 ± 0.109</td>
<td>0.249 ± 0.096</td>
</tr>
<tr>
<td>$K_L$ shower</td>
<td>1.311 ± 0.173</td>
<td>0.426 ± 0.104</td>
<td>0.642 ± 0.104</td>
<td>0.242 ± 0.091</td>
</tr>
<tr>
<td>particle ID</td>
<td>1.342 ± 0.180</td>
<td>0.414 ± 0.108</td>
<td>0.668 ± 0.107</td>
<td>0.260 ± 0.096</td>
</tr>
<tr>
<td>split-off rejection</td>
<td>1.338 ± 0.179</td>
<td>0.415 ± 0.108</td>
<td>0.667 ± 0.107</td>
<td>0.255 ± 0.095</td>
</tr>
<tr>
<td>track eff.</td>
<td>1.357 ± 0.185</td>
<td>0.446 ± 0.112</td>
<td>0.669 ± 0.110</td>
<td>0.242 ± 0.097</td>
</tr>
<tr>
<td>track resol.</td>
<td>1.317 ± 0.179</td>
<td>0.438 ± 0.108</td>
<td>0.664 ± 0.108</td>
<td>0.215 ± 0.094</td>
</tr>
<tr>
<td>split-off sim.</td>
<td>1.326 ± 0.178</td>
<td>0.432 ± 0.108</td>
<td>0.655 ± 0.106</td>
<td>0.240 ± 0.093</td>
</tr>
<tr>
<td>$K_L$ production</td>
<td>1.325 ± 0.176</td>
<td>0.431 ± 0.106</td>
<td>0.651 ± 0.105</td>
<td>0.244 ± 0.094</td>
</tr>
<tr>
<td>$\nu$ production</td>
<td>1.330 ± 0.177</td>
<td>0.432 ± 0.107</td>
<td>0.653 ± 0.105</td>
<td>0.246 ± 0.094</td>
</tr>
<tr>
<td>$\nu$ production</td>
<td>1.344 ± 0.178</td>
<td>0.425 ± 0.106</td>
<td>0.669 ± 0.106</td>
<td>0.251 ± 0.095</td>
</tr>
<tr>
<td></td>
<td>1.322 ± 0.175</td>
<td>0.439 ± 0.106</td>
<td>0.641 ± 0.104</td>
<td>0.242 ± 0.093</td>
</tr>
</tbody>
</table>

APPENDIX C: AVERAGING $|V_{ub}|$ RESULTS

In each of the $\pi \ell \nu$ and $\rho \ell \nu$ modes, we have extracted two results for $|V_{ub}|$ that are largely free from modeling assumptions: a value based on the application of LCSR-derived form factors for $q^2 < 16$ GeV$^2$, and a value based on the application of LQCD-derived form factors for $q^2 \geq 16$ GeV$^2$. We therefore have three averages to be calculated: the combination of the two results within the $\pi \ell \nu$ mode and within the $\rho \ell \nu$ mode, and the combination of the two modes. The averaging procedure should take into account, in particular, the correlations present in the systematic uncertainties in the result. This appendix describes our averaging procedure.

The statistical correlations have been taken into account in the LCSR-derived results. An evaluation of remaining statistical correlations found that they had little impact on the
TABLE XVII. Central values and statistical uncertainties for $B^0 \rightarrow \pi^- \ell^+ \nu$ branching fractions for the reference fit and for each systematic variation of the $B \rightarrow X_s \ell \nu$ background simulation input to the fit. The associated systematic uncertainties in $|V_{ub}|$ are obtained by fitting the results from the relevant set of $q^2$ intervals for each systematic study. The total branching fraction is shown as well for completeness. All results were obtained using the Ball’01 form factor for the $\pi \ell \nu$ modes and the Ball’98 form factors for the $\rho \ell \nu$ modes.

<table>
<thead>
<tr>
<th>OPE parameters</th>
<th>Hadron-ization</th>
<th>Total</th>
<th>$10^3 \times B(B^0 \rightarrow \pi^- \ell^+ \nu)$</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td>$0 &lt; q^2 &lt; 8$ GeV$^2$</td>
</tr>
<tr>
<td>nominal</td>
<td>nominal</td>
<td>1.324 ± 0.177</td>
<td>0.423 ± 0.107</td>
</tr>
<tr>
<td>nominal</td>
<td>nonres.</td>
<td>1.322 ± 0.177</td>
<td>0.431 ± 0.106</td>
</tr>
<tr>
<td>“High”</td>
<td>nonres.</td>
<td>1.311 ± 0.176</td>
<td>0.428 ± 0.106</td>
</tr>
<tr>
<td>“Low”</td>
<td>nonres.</td>
<td>1.329 ± 0.177</td>
<td>0.434 ± 0.106</td>
</tr>
</tbody>
</table>

The statistical uncertainties are uncorrelated, and are combined as

$$\sigma_{\text{stat},i} = \sqrt{\alpha \sigma_{\text{stat},i}^{16} + (1 - \alpha) \sigma_{\text{stat},i}^{16}}. \quad (C1)$$

Correlated uncertainties, such as the theoretical uncertainties, are combined as

$$\sigma_{\text{corr},i} = \alpha \sigma_{\text{corr},i}^{16} + (1 - \alpha) \sigma_{\text{corr},i}^{16}. \quad (C3)$$

For each simulation variation (labeled $i$), we perform the full analysis to obtain $|V_{ub,i}^{16}|$ and $|V_{ub,i}^{16}|$. The systematic uncertainty defined for the variation is

$$\sigma_{i} = \left| V_{ub,i}^{\text{nom}} \right|^2 - \left| \alpha |V_{ub,i}^{16}| + (1 - \alpha) |V_{ub,i}^{16}| \right|^2. \quad (C4)$$

where $|V_{ub,i}^{\text{nom}}|$ is the average resulting from Eq. (C1). This procedure preserves the systematic correlation. We combine this estimate in quadrature with the additional uncertainty contribution to allow for imperfect modeling of the correlated changes between signal efficiency and raw yield (see Sec. V).

Finally, for each value of $\alpha$ the experimental and theoretical uncertainties are combined in quadrature (taking the average theoretical uncertainty in the case of asymmetric uncertainties). We scan over $\alpha$ and choose the value that minimizes the total uncertainty.

We perform a similar procedure to combine the results from the two modes. The weights obtained individually for the different $q^2$ regions in each mode are fixed. The uncorrelated, correlated, and anticorrelated uncertainties are combined in exact analogy to the above descriptions. Taking $\beta$ as the weight of the $\pi \ell \nu$ mode in the average, we have

$$|V_{ub}\beta| = \beta |V_{ub}|^\pi + (1 - \beta) |V_{ub}|^\rho. \quad (C5)$$

For each simulation variation, the systematic estimate becomes

$$\sigma_{i} = \left| V_{ub,i}^{\text{nom}} \right|^2 - \left( \beta \left| \alpha |V_{ub,i}^{16}| + (1 - \alpha) |V_{ub,i}^{16}| \right| + (1 - \beta) \left| \alpha |V_{ub,i}^{16}| + (1 - \alpha) |V_{ub,i}^{16}| \right| \right|^2. \quad (C6)$$

These uncertainties are, as before, combined in quadrature, along with the contribution for imperfect modeling of the correlated efficiency and yield changes.

We scan over the weight $\beta$ to find the value that minimizes the overall combined uncertainty. Once again we treat the theoretical uncertainties in the $\pi \ell \nu$ and $\rho \ell \nu$ form factors as correlated in this procedure.

052001 (2000).
[51] Spurious CsI showers from hadronic interactions add linearly to $E_{miss}$, but tend to average out in the vector sum for $\vec{P}_{miss}$.
Incorrect mass assignment also smears $E_{miss}$.
[58] Particle Data Group, Hagiwara et al. [2].
[66] While the Barlow-Beeston method is used to account for statistical fluctuations in the off-resonance sample, it is imperfect when the statistics of the fit component are smaller than those of the true background component under the signal in the data. In particular, the fit will be biased by bins for which the continuum background is small and for which there was a downward Poisson fluctuation in the off-resonance sample.
[67] In principle, two-photon processes could also contribute to the continuum background. However, the dominant contribution would be from a process that scatters a beam electron or positron into the endpoint that would become the “signal” lepton. This process results in a large missing momentum down the beam pipe, and is essentially eliminated by the cos $\theta_{miss}$ requirement.